

## Kernel regression

- Nonparametric regression. Suppose that  $(X_1, Y_1), \dots, (X_n, Y_n)$  are IID data and

$$Y_i = m(X_i) + \varepsilon_i \quad (1)$$

for  $i = 1, \dots, n$ , where  $(\varepsilon_1, \dots, \varepsilon_n)$  is independent of  $(X_1, \dots, X_n)$ ,  $E(\varepsilon_1) = 0$  and  $Var(\varepsilon_1) = \sigma^2$ . The problem of interest is to estimate  $m$  based on  $(X_1, Y_1), \dots, (X_n, Y_n)$ .

- Kernel function. A kernel function  $k$  on  $(-\infty, \infty)$  usually satisfies the usual constraints:

- (a)  $k \geq 0$ .
- (b)  $\int_{-\infty}^{\infty} k(s)ds = 1$ .
- (c)  $\int_{-\infty}^{\infty} sk(s)ds = 0$ .
- (d)  $\int_{-\infty}^{\infty} s^2k(s)ds < \infty$ .

- Kernel regression estimator. Suppose that  $(X_1, \dots, X_n)$  is a random sample and  $X_i$  takes values in  $(-\infty, \infty)$  for  $i = 1, \dots, n$ . The kernel regression estimator for  $m(x)$  with kernel  $k$  and bandwidth  $h$  is

$$\hat{m}(x) = \frac{\sum_{i=1}^n Y_i k\left(\frac{x - X_i}{h}\right)}{\sum_{i=1}^n k\left(\frac{x - X_i}{h}\right)}.$$

- The estimation error  $\hat{m}(x) - m(x)$ .

$$\hat{m}(x) - m(x) = \frac{\underbrace{\frac{1}{nh} \sum_{i=1}^n (Y_i - m(x)) k\left(\frac{x - X_i}{h}\right)}_I}{\underbrace{\frac{1}{nh} \sum_{i=1}^n k\left(\frac{x - X_i}{h}\right)}_{II}}.$$

- Mean and variance of  $I$ .

$$\begin{aligned} E(I) &= E\left((nh)^{-1} \sum_{i=1}^n (Y_i - m(x)) k((x - X_i)/h)\right) \\ &= h^2 \int u^2 k(u) du \left( \frac{f(x)m''(x)}{2} + f'(x)m'(x) \right) + o(h^2), \end{aligned}$$

and

$$\begin{aligned} Var(I) &= \frac{1}{nh^2} [E((Y_1 - m(x))^2 k^2((x - X_1)/h))] \\ &\quad - \frac{h^2}{nh^2} \left[ h^2 \int u^2 k(u) du \left( \frac{f(x)m''(x)}{2} + f'(x)m'(x) \right) + o(h^2) \right]^2 \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{nh} \left[ \sigma^2 f(x) \int k^2(u) du \right] - \frac{h^4}{n} \left[ \int u^2 k(u) du \left( \frac{f(x)m''(x)}{2} + f'(x)m'(x) \right) \right]^2 \\
&\quad + o\left(\frac{1}{nh}\right) + o\left(\frac{h^4}{n}\right).
\end{aligned}$$

Thus if  $nh \rightarrow \infty$  and  $h \rightarrow 0$  as  $n \rightarrow \infty$ ,

$$\begin{aligned}
E(I^2) &= E \left( (nh)^{-1} \sum_{i=1}^n (Y_i - m(x)) k((x - X_i)/h) \right)^2 \\
&= \frac{1}{nh} \left[ \sigma^2 f(x) \int k^2(u) du \right] + h^4 \left[ \int u^2 k(u) du \left( \frac{f(x)m''(x)}{2} + f'(x)m'(x) \right) \right]^2 \\
&\quad + o\left(\frac{1}{nh}\right) + o(h^4).
\end{aligned}$$

- Mean and variance for  $II$ . Let  $f$  be the density of  $X_i$ . Suppose that  $nh \rightarrow \infty$  and  $h \rightarrow 0$  as  $n \rightarrow \infty$ , then  $E(II) - f(x) \rightarrow 0$  and  $Var(II) \rightarrow 0$  as  $n \rightarrow \infty$ , so  $II \approx f(x)$  for large  $n$  and  $E(\hat{m}(x) - m(x))^2$  can be approximated by

$$\frac{1}{nh} \left( \frac{\sigma^2 \int k^2(u) du}{f(x)} \right) + h^4 \left[ \int u^2 k(u) du \left( \frac{m''(x)}{2} + \frac{f'(x)m'(x)}{f(x)} \right) \right]^2.$$

- Kernel function on  $R^d$ . A kernel function  $k$  on  $R^d$  usually satisfies the usual constraints:

- (a)  $k \geq 0$ .
- (b)  $\int k(s) ds = 1$ .
- (c)  $\int s_i k(s_1, \dots, s_d) d(s_1, \dots, s_d) = 0$  for  $i = 1, \dots, d$ .
- (d)  $\int \|s\|^2 k(s) ds < \infty$ .

- Kernel regression estimator on  $R^d$ . Suppose that  $X_i$  takes values in  $R^d$  for  $i = 1, \dots, n$ . The kernel regression estimator for  $m(x)$  with kernel  $k$  and bandwidth  $h$  is

$$\hat{m}(x) = \frac{\sum_{i=1}^n Y_i k\left(\frac{x - X_i}{h}\right)}{\sum_{i=1}^n k\left(\frac{x - X_i}{h}\right)}.$$

- Bandwidth selection. We use leave-one-out cross validation to choose  $h$  for a given kernel  $k$ . Let  $\hat{m}_{-i,h}$  be the kernel estimator for  $m$  with bandwidth  $h$  based on  $(X_1, Y_1), \dots, (X_{i-1}, Y_{i-1}), (X_{i+1}, Y_{i+1}), \dots, (X_n, Y_n)$ . Let

$$RSSCV(h) = \sum_{i=1}^n (Y_i - \hat{m}_{-i,h}(X_i))^2.$$

Leave-one-out cross validation: choose the bandwidth  $h$  so that  $RSSCV(h)$  is minimized.

- Bandwidth selection rule(s) can be found in [1].
- An example of approximating the bias of an estimator via simulation. Suppose that  $X_1, \dots, X_n$  is a random sample from  $N(\mu, 1)$  and consider estimating  $\mu$  using the sample mean  $\bar{X}$ . The bias of  $\bar{X}$  when  $n = 50$  and  $\mu = 20$  can be approximated using simulation.

```
#generate 1000 samples of size 50 from N(20,1)
#and store the 1000 sample means in x
x <- rep(0,1000)
for (i in 1:1000){ x[i] <- mean(rnorm(50,mean=20, sd=1)) }
#compute the approximate bias (expected value for sample mean - 20 )
sum(x - 20)/1000
```

- Exercise 1. Write a function using R with the following input and output:

Input: data  $(X_1, Y_1), \dots, (X_n, Y_n)$ , kernel function  $k$ , bandwidth  $h$ , and evaluation point  $x_0$ .

Output:  $\hat{m}(x_0)$ .

You may assume  $x_0$  is one dimensional.

- Exercise 2. Consider the model in (1) with  $n = 1000$ ,  $m(x) = \sin(20x)$ , where  $X_i$  is uniformly distributed on  $[-1, 1]$  and the errors are IID  $N(0, (0.01)^2)$ . Let

$$I(x_0) = \left( (nh)^{-1} \sum_{i=1}^n (Y_i - m(x_0)) k((x_0 - X_i)/h) \right),$$

where  $k$  is the probability density function of  $N(0, 1)$ .

- Compute  $E(I(x_0))$  for  $x_0 = 0.1$  and  $h \in \{0.01, 0.005, 0.001, 0.0005\}$ . Note that

$$E(I(x_0)) = \int_{-\infty}^{\infty} (m(x_0 - hu) - m(x_0)) f(x_0 - hu) k(u) du$$

and the R command for computing  $\int_a^b g(x) dx$  is `integrate(g,a,b)`.

- Approximate  $E(I(x_0))$  for  $x_0 = 0.1$  and  $h \in \{0.01, 0.0005\}$  by IID data  $(X_1, Y_1), \dots, (X_n, Y_n)$   $10^4$  times according to (1) with the above setup.
- Exercise 3. Write a function using R with the following input and output:  
Input: data  $(X_1, Y_1), \dots, (X_n, Y_n)$ , kernel function  $k$ , a vector of bandwidths  $(h_1, \dots, h_\ell)$  and evaluation point  $x_0$ .  
Output:  $\hat{m}(x_0)$ , where the bandwidth is chosen among  $h_1, \dots, h_\ell$  using leave-one-out cross validation.

You may assume  $x_0$  is one dimensional.

## References

- [1] W. WAND AND M. JONES, *Kernel Smoothing*, Chapman & Hall, 1995.