

Independence of two random vectors

- Definition. Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are random vectors. X and Y are independent means

$$P(X \in A_1 \text{ and } Y \in A_2) = P(X \in A_1)P(Y \in A_2)$$

for all $A_1 \subset R^m$ and $A_2 \subset R^n$.

- Note. Suppose that X and Y are random vectors and X and Y are independent. Then $u(X)$ and $v(Y)$ are independent for all functions u and v .
- CDF case.

Fact 1 Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are random vectors. X and Y are independent if and only if

$$F_{X_1, \dots, X_m, Y_1, \dots, Y_n}(x_1, \dots, x_m, y_1, \dots, y_n) = F_{X_1, \dots, X_m}(x_1, \dots, x_m)F_{Y_1, \dots, Y_n}(y_1, \dots, y_n)$$

for all $(x_1, \dots, x_m) \in R^m$, $(y_1, \dots, y_n) \in R^n$.

- PDF case.

Fact 2 Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are random vectors and (X, Y) has a joint PDF. Suppose that X and Y have PDFs f_X and f_Y respectively. Define

$$g(x_1, \dots, x_m, y_1, \dots, y_n) = f_X(x_1, \dots, x_m)f_Y(y_1, \dots, y_n)$$

for $(x_1, \dots, x_m) \in R^m$ and $(y_1, \dots, y_n) \in R^n$, Then g can be a PDF of (X, Y) if and only if X and Y are independent.

- PMF case.

Fact 3 Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are discrete random vectors and (X, Y) has a joint PMF $p_{X,Y}$. Let p_X and p_Y be the PMFs of X and Y respectively. Then X and Y are independent if and only if

$$p_{X,Y}(x_1, \dots, x_m, y_1, \dots, y_n) = p_X(x_1, \dots, x_m)p_Y(y_1, \dots, y_n)$$

for $(x_1, \dots, x_m) \in R^m$ and $(y_1, \dots, y_n) \in R^n$.

- General case.

Fact 4 Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are random vectors. X and Y are independent if and only if $\langle a, X \rangle$ and $\langle b, Y \rangle$ are independent for all $a \in R^m$, $b \in R^n$.

For $x = (x_1, \dots, x_n)$ and $y = (y_1, \dots, y_n)$, the notation $\langle x, y \rangle$ means $\sum_{i=1}^n x_i y_i$, the inner product of x and y .

- MGF case.

Fact 5 Suppose that $X = (X_1, \dots, X_m)$ and $Y = (Y_1, \dots, Y_n)$ are random vectors with joint MGF $M_{X,Y}$ and $M_{X,Y}(s_1, \dots, s_m, t_1, \dots, t_n) < \infty$ for (s_1, \dots, s_m) and (t_1, \dots, t_n) such that $\max\{|s_1|, \dots, |s_m|, |t_1|, \dots, |t_n|\} < \delta$ for some $\delta > 0$. Then X and Y are independent if and only if

$$M_{X,Y}(s_1, \dots, s_m, t_1, \dots, t_n) = M_X(s_1, \dots, s_m)M_Y(t_1, \dots, t_n)$$

for (s_1, \dots, s_m) and (t_1, \dots, t_n) such that $\max\{|s_1|, \dots, |s_m|, |t_1|, \dots, |t_n|\} < \delta$, where M_X and M_Y are the MGFs of X and Y respectively.

Note that

$$M_X(s_1, \dots, s_m) = M_{X,Y}(s_1, \dots, s_m, 0, \dots, 0)$$

for $(s_1, \dots, s_m) \in R^m$ and

$$M_Y(t_1, \dots, t_n) = M_{X,Y}(0, \dots, 0, t_1, \dots, t_n)$$

for $(t_1, \dots, t_n) \in R^n$.

- Example 1. Suppose that X_1, X_2, X_3 are independent random variables and the distribution of each X_i is $N(\mu, \sigma^2)$, where μ and σ are constants and $\sigma > 0$. Let $\bar{X} = (X_1 + X_2 + X_3)/3$ and $Y = (X_1 - \bar{X}, X_2 - \bar{X}, X_3 - \bar{X})^T$. Show that \bar{X} and Y are independent.

Sol. We will first find the joint MGF of \bar{X} and Y . From Example 4(a) in the handout “Independent random variables”,

$$E(e^{wX_1}) = e^{\mu w + 0.5\sigma^2 w^2}$$

for $w \in R$. Let $M(w) = E(e^{wX_1})$ for $w \in R$. For $s \in R$ and $t = (t_1, t_2, t_3)^T \in R^3$,

$$\begin{aligned} E(e^{s\bar{X} + t^T Y}) &= E \exp \left(s\bar{X} + \sum_{i=1}^3 t_i (X_i - \bar{X}) \right) \\ &= E \exp \left(\sum_{i=1}^3 t_i X_i + (s - (t_1 + t_2 + t_3))\bar{X} \right) \\ &= E e^{a_1 X_1 + a_2 X_2 + a_3 X_3}, \end{aligned}$$

where $a_i = t_i + (s - (t_1 + t_2 + t_3))/3$ for $i = 1, 2, 3$. Since X_1, X_2, X_3 are independent and have the same distribution $N(\mu, \sigma^2)$, we have

$$E(e^{s\bar{X} + t^T Y}) = M(a_1)M(a_2)M(a_3)$$

$$\begin{aligned}
&= \prod_{i=1}^3 M(t_i + (s - (t_1 + t_2 + t_3))/3) \\
&= \exp \left(\sum_{i=1}^3 \left[\mu \left(t_i + \frac{s - 3\bar{t}}{3} \right) + 0.5\sigma^2 \left(t_i + \frac{s - 3\bar{t}}{3} \right)^2 \right] \right),
\end{aligned}$$

where $\bar{t} = (t_1 + t_2 + t_3)/3$. It can be shown that

$$\begin{aligned}
&\sum_{i=1}^3 \left[\mu \left(t_i + \frac{s - 3\bar{t}}{3} \right) + 0.5\sigma^2 \left(t_i + \frac{s - 3\bar{t}}{3} \right)^2 \right] \\
&= \mu \cdot s + \frac{\sigma^2 s^2}{6} + 0.5\sigma^2 \sum_{i=1}^3 (t_i - \bar{t})^2,
\end{aligned} \tag{1}$$

where $\bar{t} = (t_1 + t_2 + t_3)/3$. Therefore, for $s \in R$ and $t = (t_1, t_2, t_3)^T \in R^3$,

$$E(e^{s\bar{X} + t^T Y}) = M_1(s)M_2(t),$$

where

$$M_1(s) = \exp \left(\mu \cdot s + \frac{\sigma^2 s^2}{6} \right)$$

and

$$\begin{aligned}
M_2(t) &= \exp \left(0.5\sigma^2 \sum_{i=1}^3 (t_i - \bar{t})^2 \right) \\
&= \exp \left(0.5\sigma^2 \sum_{i=1}^3 (t_i - (t_1 + t_2 + t_3)/3)^2 \right).
\end{aligned}$$

It is clear that

$$E(e^{s\bar{X} + t^T Y}) \Big|_{t=(0,0,0)^T} = M_1(s)$$

and

$$E(e^{s\bar{X} + t^T Y}) \Big|_{s=0} = M_2(t),$$

so

$$E(e^{s\bar{X} + t^T Y}) = \left(E(e^{s\bar{X} + t^T Y}) \Big|_{t=(0,0,0)^T} \right) \left(E(e^{s\bar{X} + t^T Y}) \Big|_{s=0} \right)$$

for $s \in R$, $t \in R^3$, which implies that \bar{X} and Y are independent.

- IID= Independent and Identically Distributed (獨立同分布). Suppose that random variables X_1, \dots, X_n are independent and identically distributed, then we say that X_1, \dots, X_n are IID.
 - In Example 1, X_1, X_2, X_3 are IID.